

Psychometric Qualities of a Core Set to Ascertain the Functional Profile of Portuguese Elderly Citizens

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Abstract. Objectives: This paper describes the psychometric qualities of a core set composed initially of 31 codes and extracted from International Classification of Functioning, Disability and Health, to ascertain the Functional Profile of Portuguese Elderly Citizens, residing in their own home or at a family or friends' home. **Methods:** Cross-sectional, descriptive study, with a final sample totaled 351 elderlies. Data collected by health professionals in the participants' houses, using the Elderly Nursing Core Set questionnaire. **Results:** The recommendation of the construct to the EFA was "excellent". Regarding reliability, the construct revealed factorial reliability. In terms of validity, the construct presented factorial validity and convergent validity, although failing regarding discriminant validity. **Discussion:** Comparing psychometric qualities between the original Elderly Nursing Core Set previously applied to institutionalized citizens in relation to the one presented in this paper (citizens residing in their own home or at a family or friends' home), lead to five latent factors and differences between functional profiles. More than half of the citizens are married and almost half of the sample never went to school, thus revealing an important aspect characterizing a lower literacy level of the citizens involved in this research.

Conclusions: The work based on Core Sets extracted from the International Classification of Functioning, Disability and Health, delineated to assess the nursing care needs and/or the outcomes of nursing interventions of citizens aged 65 years old or older, will be an ongoing process that will lead to the promotion of an Healthy Ageing and functional ability, as stated by World Health Organization.

Keywords: Ageing · Elderly residing in the community · Functionality profile assessment · Confirmatory Factor Analysis

1 Introduction

Demographic data have been shown that Portugal is one of the European countries presenting one of the highest ageing index (153.2 elderly per 100 young), with this ageing

phenomena being an ongoing process that will probably last the next fifty or sixty years [1]. This change in the age profile in Portugal (like in other countries of the world) is mainly characterized by two simultaneously processes: (i) the reduction of people with age less than 15 years (a decrease in birth rate and an increase in the emigration of Portuguese young people [2]) along with; (ii) the raise of the average life expectancy that lead to the increase of people with 65 years and more. According to the Statistics Portugal Institute (Instituto Nacional de Estatística – INE) the life expectancy was 80.80 years at birth for the entire Portuguese population, between 2016 and 2018. On the other hand, life expectancy at age 65 attained 19.49 years, also for the entire population, but distributed between man and women as follows: (i) a man aged 65 years could expect to live another 17.58 years; (ii) and woman aged 65 another 20.88 years [3].

However, when introducing “Health” in the context of life expectancy, regarding citizens with aged 65 years, it comes up the concept of “Healthy life years at 65 years”. This concept introduces the number of remaining years that a person aged 65 is still expected to live in a healthy condition, i.e., citizens living with the absence of limitations in functioning/disability during a certain period of time. According to the most recent statistics data (year 2017), the healthy life years at 65 for the Portuguese population, by sex, is as follows: (i) 7.9 years for mean; (ii) 6.7 years for mean; whose data puts Portugal almost among with last group of European countries whose citizens that live with less number of Healthy life years [4]. However, despite the aforementioned healthy life years at 65 years, there are citizens who already live with the presence of disabilities after this age, or even earlier.

A recent international report was done, to assess the Portuguese health system in relation to its performance on key challenges and opportunities in the post-financial crisis recovery period, involving the European Observatory on Health Systems and Policies, and the Regional Office for Europe of the World Health Organization (WHO), as well as the Health Portuguese authority (Serviço Nacional de Saúde – SNS). One of the reviewed topics was “Multimorbidity”, and the authors of the report state that most adults attending primary health care have more than one chronic condition. Data provided on this report (referred to 2016 and designated by “Burden of Chronic Disease”) shows that 41% users of Portuguese Healthcare System presented multimorbidity (they have equal or more than two chronic diseases (who 53.7% of them presented four or more chronic diseases), with only 18% presented one, thus resulting on 41% of citizens free from any chronic disease [5]. Moreover, multimorbidity may cause serious implications when health personnel need to provide care delivery to citizens in this health condition, which may become a challenge task, because the treatment of one chronic disease will be in the context of other chronic conditions. Regarding age, the same report states that major percentage of citizens with chronic diseases occur between the following age groups: (i) ages between [60–70], 70% for man and 75% for women; (ii) and ages between [80–90], 81% for man and 83% for women; (iii) followed by a decrease regarding the next age group (ages between [90–150]), more pronounced in man when compared to women, with a medium percentage of 72% of citizens presenting chronic disease [5].

As citizens are getting older, they become increasingly fragile, presenting functional impairments, multimorbidity (already shown based on the aforementioned data), and a significant prevalence of chronic conditions that easily decompensate them, leading

to progressive losses, as envisioned by the various biological, psychological and social theories, frequently originating acute health care situations [6–8]. Due to the complexity and heterogeneity of individual aging, the level of functionality may vary distinctly from person to person. Studying the individual aging by integrating a care model in continuity and proximity, allowing the elderlies and family caregivers to monitor and manage their health at home, always under the supervision of health professionals, may result in management of various chronic conditions (multimorbidity) and to provide an appropriated “safety net” before occurring an “health crisis”. To achieve this goal, i.e., to promote the quality of life related to the elderly’s health, by requalifying their potential and allowing them to live with more independence and autonomy, it is essential: (i) to evaluate his/her functional capacity in order to identify their disabilities; (ii) identify his/her appropriate self-care behavior that allows diagnosing, planning and assessing the necessary preventive nursing care needs, assuming that the demands of nursing care are high among older adults [9, 10]. Additionally, according to other authors, citizens with multimorbidity have better health outcomes when they benefit from adequate health interventions, structured from previous assessments of their functional level [8, 11], allowing us to get the “big picture” in regard to the presence of disabilities and the overall health state of a person, around several components of his/her life, such as physical, psychological, social and environmental.

To achieve this goal, i.e., to classify the degree of functioning, the WHO developed the International Classification of Functioning, Disability and Health (ICF), which is the framework for measuring health and disability at both individual and population levels [12]. However, the full ICF taxonomy encompasses an extremely large number of elements. To develop a more manageable means of assessing functioning, several core sets (sets of ICF codes) have been developed [13].

In Portugal, the authors César Fonseca et al. have been doing a very important and pioneer work on studying indicators of disability, and more particularly regarding limitations in activities and limitations in the functional capacity, targeting people aged 65 years and older. This important work resulted in grouping sets of ICF codes (core set), namely the “Elderly Nursing Core Set”, aiming to classify the degree of elderlies functioning, which is also capable of even establishing elderly “nursing care needs” [10]. However, as mentioned earlier, their target was people already institutionalized.

Therefore, differing from the work developed by César Fonseca et al., the goal of this research is to study the Psychometric Qualities of the ENCS to ascertain functional profiles of Portuguese elderly citizens living in a rural area at the main Portuguese territory (people residing in their own homes or at a family or friends’ home, i.e., in the community, thus not institutionalized). Other aims of the proposed work are as follows: (i) extract (through an Exploratory Factor Analysis– EFA) a set of latent factors that explained the relational structure of items applied to elderly resident citizens; (ii) validate the model extracted from the EFA through a Confirmatory Factor Analysis (CFA), which is not provided in [10]; (iii) a comparison between average functional scores of the entire sample, based on ENCS strategy and the computations of those using the extracted (*factor score weights*) *fsw* resulting from CFA (also not provided in [10]).

Finally, this research will contribute to the lack of population-based studies on socioeconomic, demographic, and health characteristics of the adult population living in rural areas [14].

2 Methods

2.1 Subjects

This research work involved a sample of citizens with age of 65 and more (elderlies) residing in a rural area namely Baixo Alentejo – BA, located at middle southeast of Portugal main territory that makes border with Spanish region namely Extremadura. 468 participants were selected by stratified random sampling of all 32893 citizens registered in the database of the Health Local Unit of Baixo Alentejo Region (Unidade Local de Saúde do Baixo Alentejo – ULSBA [15]). The sample was stratified by sex (male and female) and by aged group (65–74, 75–84 and 85 and more years old). Data were collected by health professionals in the participants' houses, using the Elderly Nursing Core Set questionnaire developed in [10], between January 2016 and April 2017.

The inclusion criteria rules (only comprising citizens aged 65 or older) were, cumulatively: (i) individuals interested in participating in the study; (ii) residing at the BA region in their own homes or at a family or friends' home; and (iii) those who were able to make decisions, even if sick or hospitalized. The final sample totaled 351 elderlies, those who answered the ENCS questionnaire fully and correctly, and signed the respective informed consent, as well as fulfilled all the stages of the inclusion criteria (response rate of 75%). More details about the construction, the main characteristics and the scale used by the codes (items) included in the ENCS can easily be found in [10].

2.2 Statistical Methods

The list of 31 ICF codes identified by César Fonseca et al. is shown in Table 1 (extracted from [16]).

Table 1. ICF codes included in ENCS [16].

ICF descriptors [17]	ICF codes [18]
Body functions	b110, b114, b140, b144, b152, b164, b280, b420, b440, b525
Body structures	s810
Activities and participation	d230, d310, d330, d350, d410, d415, d445, d450, d465, d510, d520, d530, d540, d550, d560, d760
Environment	e310, e320, e340, e355

The first working stage started by an Exploratory Factor Analysis (EFA), including the 31 codes listed in Table 1, in order to find the number of latent factors that explained the relational structure of the items, using IBM SPSS Statistics version 24.0.0 (IBM,

Armonk, NY), as described in Marôco [19]. After that, the resulting factorial structure was processed by a Confirmatory Factor Analysis using the software AMOS (v.24, SPSS, an IBM company, Chicago, IL), in order to obtain the respective factorial validity of the resulting factorial structure, as suggested in Marôco [20]. The third stage encompassed a CFA with a second order factor, in order to infer each *fsw* ICF codes, which can be used to compute the functional profiles scores, and compared them with those computed adopting an unitary weight for each item, as proposed by the authors in [10].

The final task comprised a short descriptive analysis, to describe the biological and sociodemographic variables of the sample data, using absolute and relative frequencies.

3 Results

3.1 Exploratory Factor Analysis

After checking some few deviations of normality assumption of the 31 items (codes ICF), based on the analysis of skewness (Sk) and kurtosis (Ku), notably using thresholds of $|Sk| < 3$ and $|Ku| < 10$ (as suggested by Marôco in [19]), the EFA analysis was based on the Spearman's correlation matrix. Results can be shown in Table 2.

Results presented in Table 2 shows that the obtained factorial structure exhibit five latent factors (after iteratively removal of codes b280, b420, b440, b445, b525, and s810, which started with the initial number of eight latent factors), following the rule of "eigenvalue greater than 1". Almost 100% of the 25 factor loadings are equal or greater than 0.5, while four latent factors presented mean values of the respective factor loadings equal or greater than 0.7, except for factor "F5" due to some lower factor loadings, but all greater than 0.5. Regarding the communalities, almost all were large (only 20% present values less than 0.5, but none of them are below than 0.3, which can be a threshold considered appropriate for this type of construct). The total variance explained by the model is higher than 60% (64.3%), with the "F1" latent factor explaining more than 50% of this variance. Finally, the recommendation to the EFA was considered as "Excellent", according to the Kaiser-Meyer-Olkin adequacy measure ($KMO = 0.909$). Moreover, regarding some adjustment indexes to evaluate the quality of the model adjustment, like the goodness-of-fit and the adjusted goodness-of-fit indices, they present relative poor, ($GFI = 0.630$ and $AGFI = 0.350$), which may be related to the relative low factor loadings

o "F5", although a "very good" $RMSR^*$ of 0.05 was achieved

Regarding the contents of the ICF codes (items of EFA), we produced the following

thematic classification of the five Latent factors, based on [17, 18]: (i) "F1" = "Self-care in activities of daily living"; (ii) "F2" = "Self-care in fundamental human needs"; (iii) "F3" = "Mental functions"; (iv) "F4" = "Communication"; (v) "F5" = "Support and relationships", with some similarity to the one proposed by Fonseca et al. [10].

Table 2. EFA results, leading to five retained latent factors, respective factor weights and communalities.

ICF codes (items)	Latent factors ^a					Communalities
	F1	F2	F3	F4	F5	
d450	.807	–	–	–	–	.716
d410	.804	–	–	–	–	.747
d465	.763	–	–	–	–	.666
d415	.761	–	–	–	–	.646
d230	.753	–	–	–	–	.703
d510	.651	–	–	–	–	.720
d520	.649	–	–	–	–	.714
d540	.568	–	–	–	–	.711
Mean value	.720	–	–	–	–	–
d550	–	.808	–	–	–	.791
d560	–	.781	–	–	–	.772
d530	–	.492	–	–	–	.508
Mean value	–	.694	–	–	–	–
b114	–	–	.816	–	–	.792
b110	–	–	.809	–	–	.795
b140	–	–	.795	–	–	.761
b152	–	–	.587	–	–	.489
b144	–	–	.570	–	–	.523
b164	–	–	.495	–	–	.595
Mean value	–	–	.679	–	–	–
d350	–	–	–	.799	–	.821
d330	–	–	–	.741	–	.751
d310	–	–	–	.662	–	.713
Mean Value	–	–	–	.734	–	–
e310	–	–	–	–	.718	.548
e320	–	–	–	–	.644	.469
e355	–	–	–	–	.533	.336
e340	–	–	–	–	.515	.390
d760	–	–	–	–	.504	.403
Mean value	–	–	–	–	.583	.643
<i>Eigenvalues</i>	9.749	1.606	2.156	1.438	1.129	–
<i>Variance explained</i>	39.0%	6.4%	8.6%	5.8%	4.5%	–
<i>Cronbach's alfa</i> ^b	$\alpha = 0.924$ (Very good)	$\alpha = 0.779$ (Reasonable)	$\alpha = 0.848$ (Good)	$\alpha = 0.853$ (Good)	$\alpha = 0.580$ (Almost unallowable)	–

^aExtraction Method: Principal Component Analysis.

Rotation Method: Varimax with Kaiser Normalization;

Rotation converged in 7 iterations.

^bQualitative classification adopted from [19].

3.2 Confirmatory Factor Analysis

The second stage of this researched consisted on a CFA, using the factorial structure obtained from EFA (see Table 2), as a testable model. The computation started based on a first order model as shown in Fig. 1. The names included in the symbols representing the five latent factors are abbreviations of the thematic names that were given in the above section, notably: (i) “Self-care in activities of daily living” – “SC-ADL”; (ii) “Mental Functions” – “MF”; (iii) “Self-care in fundamental human needs” – “SC-FHN”; (iv) “Communication” – “COM”; (v) “Support and relationships” – “SR”. The goodness of fit of the adjusted model can be assessed by reading the respective adjustment indexes listed at the top of Fig. 1. All the thematic names given to each latent factor were done in respect to the contents of ICF codes description, the name of ICF groups of codes [17, 18], as well as including the recommendations described in [10].

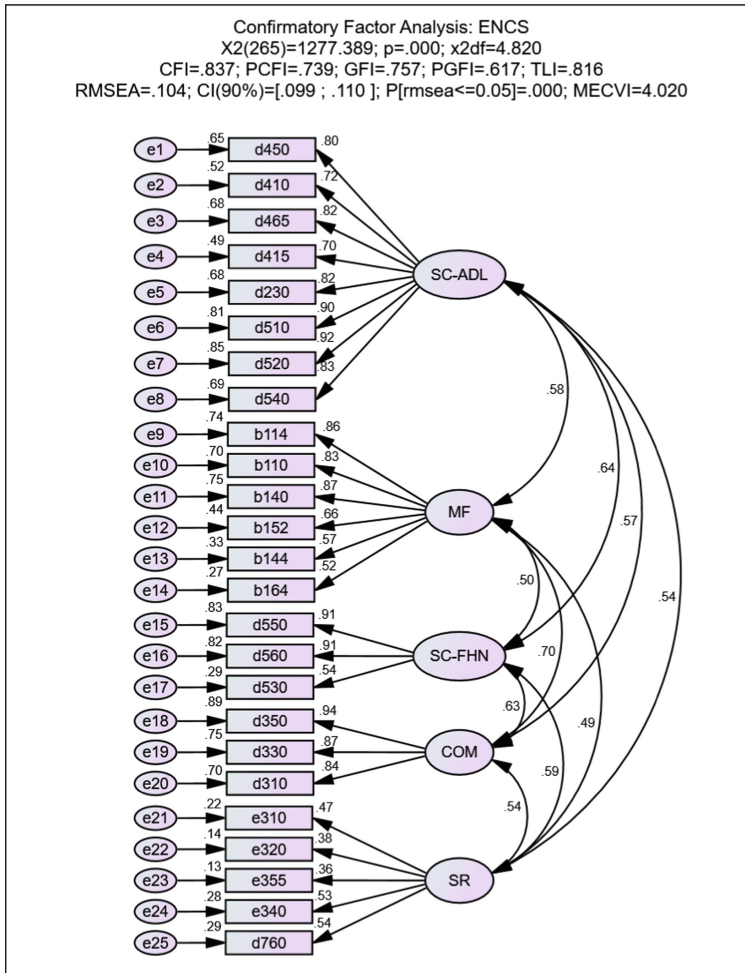


Fig. 1. Initial CFA model including the 25 items extracted from the EFA.

As can be seen, the model fit is poor: (i) Chi-squared statistics divided by the degrees of freedom (χ^2/df) is almost equal to 5; (ii) Comparative fit index (*CFI*) stays between [0.8; 0.9]; (iii) Goodness-of-fit index (*GFI*) is less than 0.8; (iv) Root Mean Square Error of Approximation (RMSEA) is greater than 0.1, among other issues (see Table 4.1 in [20] for the qualitative classification).

However, to achieve a better fit, the errors of observed variables (referred from “e1” until “e25”), were correlated, based on the modification indexes (MIs > 11) as suggested by Marôco in [20]. As can be seen, the Mean Expected Cross-Validation Index (*MECVI*) shown at the top of Fig. 2 (*MECVI* = 2.897) is lower than the previous model (*MECVI* = 4.020), which allows inferring that model in Fig. 2 presents a better global fit when compared to the model shown in Fig. 1.

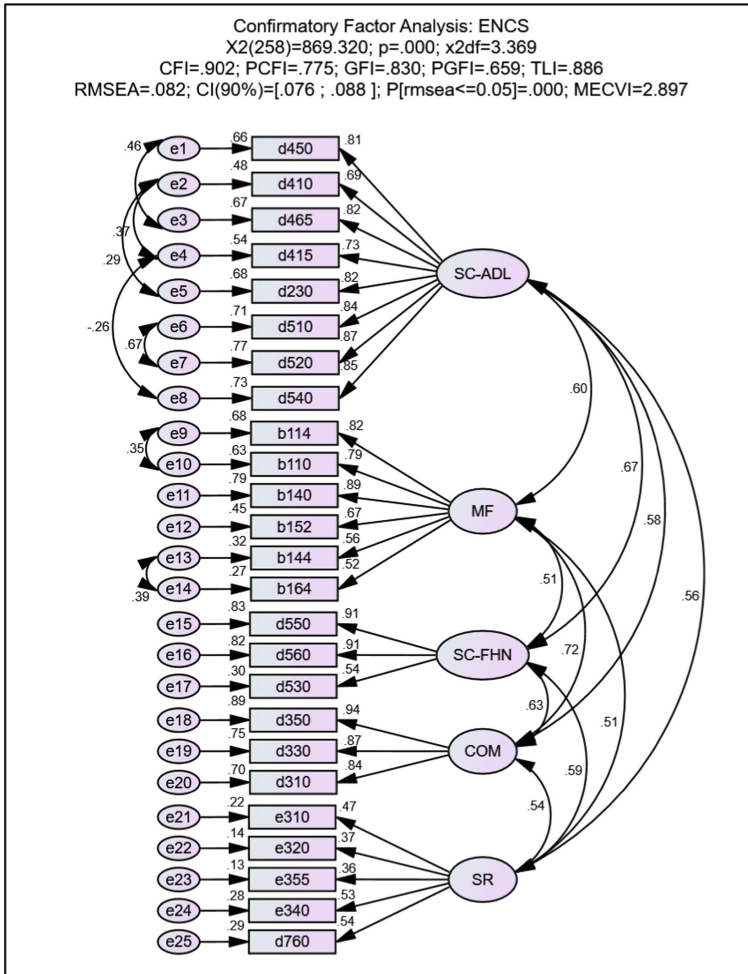


Fig. 2. Adjusted model based on the modification indexes (MI > 11, this threshold suggested by Marôco in [20])

The quality of the adjusted model can be now classified as reasonable: (i) Chi-squared statistics divided by the degrees of freedom (χ^2/df) is almost equal to 3; (ii) Comparative fit index (*CFI*) is “good”, staying between [0.9; 0.95]; (iii) Goodness-of-fit index (*GFI*) is “favorable”, staying between [0.8; 0.9]; (iv) Root Mean Square Error of Approximation (RMSEA) is “acceptable”, staying between [0.05; 0.1]; Tucker-Lewis Index (*TLI*) is “favorable”, staying between [0.8; 0.9], (see Table 4.1 in [20] for the qualitative classifications used here). Regarding the model adjustment based on the MI values, it is important to refer that the process must be grounded on some theoretical aspects of the model, and not only correlate those errors because the involved MI values are high. For example, we assume that some of the items present some similarity in their formulation, and their contents are some way related, as an example: “d410 Changing basic body position” (brushing teeth, shaving, grooming, etc.) and “d415 Maintaining a body position” (bathing, drying, washing hands, etc.). As can be seen, these both codes refer the same body positions, like Squatting, Kneeling, Sitting, Standing, Bending, which may confuse respondents. The items individual reliability was measured by the respective standardized factor loadings (λ), with only two presenting values lower than 0.4 and a third almost equal to 0.5, which allows us to assume a “Favorable” factorial validity of the construct. The construct reliability (internal consistency) was evaluated through the Cronbach’s Alfa (α_c) and composite reliability (CR), whose values are listed in Table 3 and almost greater than 0.7, as recommended in Marôco [20]. However, according to Hair et al. [21], values less than 0.7 may be accepted in case of exploratory research, which is the case here, thus we also assume a “favorable” reliability of the construct. Finally, regarding the construct validity (CV), we started by assuming a factorial validity of the construct, since it was checked that the items effectively conceive the “big picture” that is actually being measured by the specific latent factors. In relation to the convergent validity of the construct, this characteristic was evaluated by the average variance extracted of the CFA model (AVE) [20]. The results in Table 3 shows that only “SR” factor present a “fair” convergent validity (AVE values must be greater than 0.5, as suggested in Marôco [20]) which is due to the fact that this latent factor includes three standardized regression factor loadings lower than 0.5 (see Fig. 2). However, it is possible to foresee an almost overall “favorable” convergent validity of the adjusted CFA model.

Table 3. List of α_c , CR and AVE index values.

Latent factors	α_c	CR	AVE
SC-ADL	0.924	0.938	0.654
MF	0.848	0.863	0.521
SC-FHN	0.779	0.841	0.649
COM	0.853	0.914	0.780
SR	0.580	0.564	0.210

As an alternative solution to some high correlations between latent factors shown in model of Fig. 2 (all with $p < 0.001$), we suggest a hierarchical model positioning an

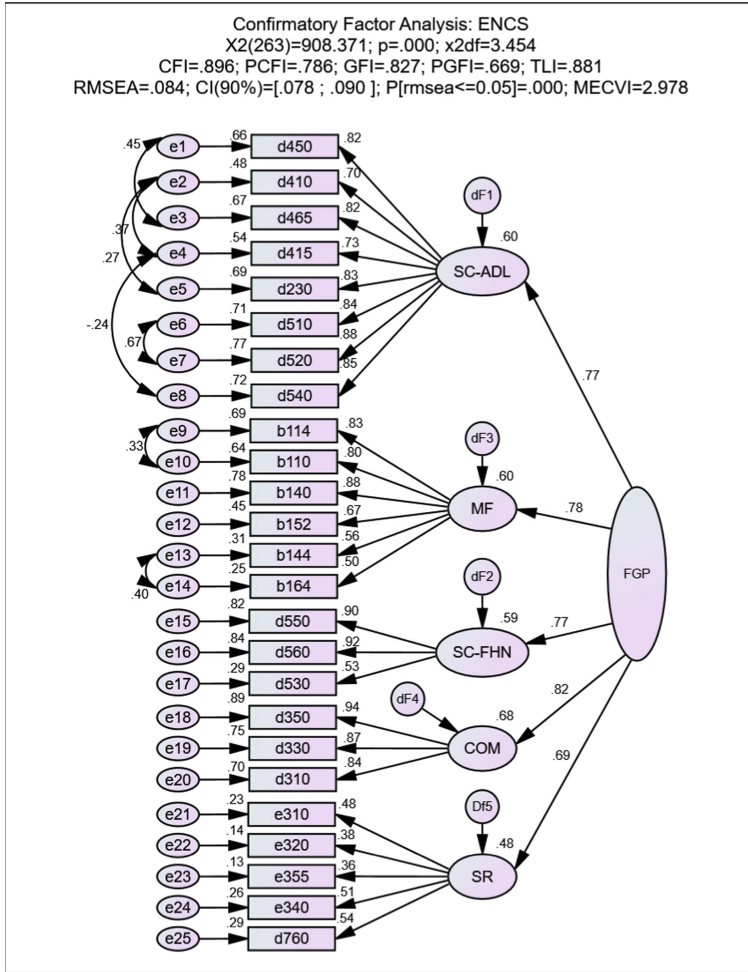


Fig. 3. Fitted second order factorial model adopting a FGP latent factor, after correlating the measurement errors of facets whose MI suggested their correlation (adopted MI > 11).

high order factor called Functioning General Profile - FGP (second order latent factor) as shown in Fig. 3, as recommended for this type of models in Marôco [20].

As can be seen by the values at the top of Fig. 3, this second order model shows a reasonable fit indexes (similar to the ones presented in top of Fig. 2), with the second factor order being the FGP measure, expressed through the various items (observed variables) and associated with each of the five latent factors. In respect to the correlations between FGP and the five latent factors, all were somehow high and all are statistical highly significant ($p < 0.001$): (i) $\rho_{SC-ADL} = .77$; (ii) $\rho_{MF} = .78$; (iii) $\rho_{SC-FHN} = .77$; (iv) $\rho_{COM} = .82$; (v) $\rho_{SR} = .69$. Regarding FC, α_c and AVE index values for FGP latent factor, respective very good values were obtained: 0.878, 0.921 and 0.591, respectively.

Based on the CFA results, a factor score weight (fsw) is extracted for each item (ICF code), which allows computing the scores of those six latent factors (SC-ADL, SC-FHN, MF, COM, SR and FGP), for each citizen individually. According to the authors of the original ENCS [10], the score of each latent factor is computed as the average value of the items (ICF codes) that are included in the respective latent factor, thus using a fsw value of “1” for each item. However, in this paper we propose the computation of individual’s functional profile scores of each latent factor based on the fsw available for each item. Table 4 shows the adjusted fsw and how to compute the new scores of each latent factor (profiles).

Table 4. Proposed formulae used to compute the individual’s functional profile scores based on the fsw extracted from the CFA second order model shown in Fig. 3.

Latent factor	Formulae ($fsw \times$ individual responses to the ICF codes)
SC-ADL	$0.104 \times d450 + 0.001 \times d410 + 0.104 \times d465 + 0.156 \times d415 + 0.139 \times d230 + 0.066 \times d510 + 0.172 \times d520 + 0.258 \times d540$
SC-FHN	$0.406 \times d550 + 0.557 \times d560 + 0.037 \times d530$
MF	$0.241 \times b114 + 0.176 \times b110 + 0.416 \times b140 + 0.104 \times b152 + 0.041 \times b144 + 0.022 \times b164$
COM	$0.556 \times d350 + 0.285 \times d330 + 0.159 \times d310$
SR	$0.159 \times e310 + 0.093 \times e320 + 0.125 \times e355 + 0.289 \times e340 + 0.334 \times d760$
FGP	$0.016 \times d450 + 0.000 \times d410 + 0.016 \times d465 + 0.024 \times d415 + 0.022 \times d230 + 0.010 \times d510 + 0.027 \times d520 + 0.040 \times d540 + 0.113 \times d550 + 0.154 \times d560 + 0.010 \times d530 + 0.047 \times b114 + 0.034 \times b110 + 0.081 \times b140 + 0.020 \times b152 + 0.007 \times b144 + 0.004 \times b164 + 0.154 \times d350 + 0.079 \times d330 + 0.044 \times d310 + 0.016 \times e310 + 0.009 \times e320 + 0.012 \times e355 + 0.029 \times e340 + 0.033 \times d760$

Based on the results listed in Table 4, it was possible to compute the average functional profiles for the entire sample, whose results are provided in Fig. 4. The “Mean*” values represent the scores computed based on the formulation proposed in [10], while “MeanFSW” values symbolize those scores computed using the formulae listed in Table 4. The remaining values in Fig. 4 represent: (i) “MaxDiff”, the maximum positive differences at individual level (adopting “Mean*”-“MeanFSW”); (ii) “Min-Diff”, the minimum negative difference for the same minus operation, also at individual level; (iii) “StdDiff”, the standard deviation between all individual differences, notably “MaxDiff” and “MinDiff”. All the items were in Likert scale (1–5) and scores (ranging from 0 to 100%) were computed based on the following equation:

$$Functional\ score = 25 \times itens\ response - 25. \quad (1)$$

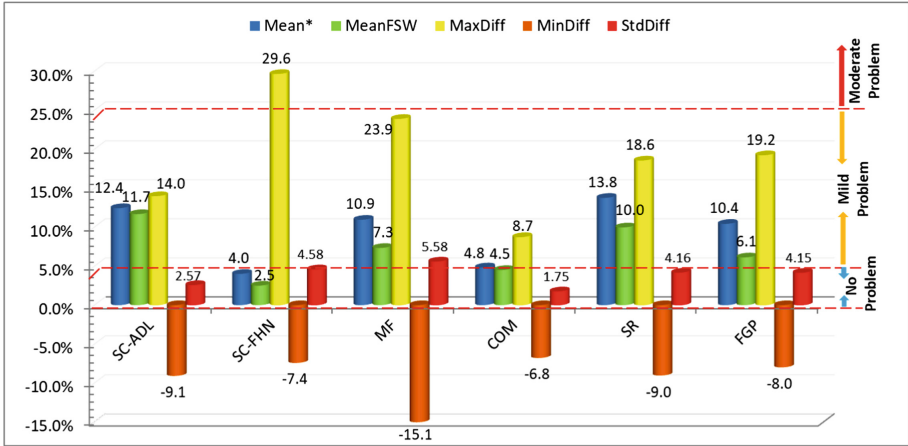


Fig. 4. Average functional profiles scores (for the entire sample for each of the six latent factors), computes using the *fsw* extracted from the adjusted second order CFA model.

3.3 Descriptive Statistics

A short descriptive statistics of the sample data, processed under the scope of this research, is also provided in Table 5. The table lists the biological and sociodemographic characteristics of the 351 respondents residing in the region of BA.

Table 5. Biological and sociodemographic characteristics of the sample data.

Biological and sociodemographic variables	<i>N</i>	%
<i>Gender:</i>	–	–
Male	163	46.4
Female	188	53.6
<i>Age group:</i>	–	–
65–74	132	37.6
75–84	135	38.5
85 and higher	84	23.9
<i>Marital status:</i>	–	–
Single	27	7.7
Married	206	58.7
Divorced	4	1.1
Widowed	114	32.5
<i>Educational level:</i>	–	–
Does not know how to read or write	104	29.6
Knows how to read and write	59	16.8
1 st –4 th grade	165	47.0
More education	23	6.6

4 Discussion

Comparing the work developed in [10] with the one delineated in this paper, both show an “Excellent” recommendation regarding to the EFA, since the Kaiser-Mayer-Olkin adequacy measures are between the range] 0.9; 1.0] [20]: 0.963 and 0.909, respectively. Moreover, the work in [10] was able to extract four latent factors, while the one presented here is composed of five latent factors, whose the major difference is related to self-care: (i) one latent factor designated by “Self-Care” in [10]; (ii) two latent factors in the present proposal, namely “Self-care in activities of daily living” (composed by eight ICF codes) and “Self-care in fundamental human needs” (composed by three ICF codes). In relation to this topic, the present proposal may provide a little more detail regarding the citizens Self-care needs. Given the results of Bartlett’s test of sphericity ($p < 0.001$), we rejected the null hypothesis and concluded that the variables of both constructs were significantly correlated. In addition, we established that the 25 indicators included in this component exhibit excellent internal consistency as also in [10].

Regarding the CFA results, adjustment indexes shows a very favorable adjusted CFA model, with the “SR” latent factor exhibiting the lowest reliability, motivated by three lower loadings, ranging from 0.5 and 0.3, with this lower threshold (0.3) being considered appropriate for this type of construct. Important issues regarding the validation of the construct based on CFA, reveal that the items achieved a very favorable individual reliability, and the construct reliability as also reached, although the “SR” latent factor was the one that presented a lower contribution for this validation issue, although values below 0.7 and greater than 0.5 may be accepted in case of exploratory research, like the one developed here, according to Hair et al. [21]. In relation to the construct validity, the factorial validity of the construct was also reached, since we checked that the items effectively conceive the “big picture” that is actually being measured by the specific latent factors. Regarding the convergent validity, only the “SR” latent factor presented a value less than 0.5, which can be explained some of the lower loading factors of the respective items, whose mean values was 0.46 $((0.47 + 0.37 + 0.36 + 0.53 + 0.54)/5)$. Finally the construct fails in terms of discriminant validity, because correlations between first-order factors were higher, associated to lower AVE values. In this point, a comparison to the work developed by César et al. it was not possible, because no CFA is available in [10].

Regarding the functional profiles scores calculated based on the *fsw* extracted from CFA model shown in Fig. 3 and those computed according to the details provided in [10], the full scenario is depicted in Table 5. The analysis of the results shows that the global mean scores are almost similar for the “SC-ADL”, “SC-FHN” and “COM” latent factors, and more pronounced for “MF”, “SR” and “FGP”, although always higher for “Mean*” than “MeanFSW”. Almost all the global means reach a “Mild functional profile”, except for “SC-FHN” (corresponding to “No Problem” functional profile). The most relevant aspect found is when making an individual analysis of the scores of each citizen, because there are differences in the results, which are higher positive differences for “SC-FHN”, “MF”, “SR” and “FGP”, when compared to “SC-ADL” and “COM”, and major negative differences in “MF”. Those differences were the maximum found (either positive or negative), among all citizens composing the sample. We think that the differences between scores found must be carefully analyzed and it should be done by a

group of health experts. Again, comparison at this level to the work developed by César et al. it was not possible, because no CFA is available in [10].

Finally, biological and sociodemographic characteristics of the sample show the presence of more females than males, the group of citizens presenting age a more advanced age as a relevant proportion in comparison to the other two. More than half of the citizens are married and almost half of the sample ($29.6\% + 16.8\% = 46.4\%$) never went to school, thus revealing an important aspect characterizing a lower literacy level of the citizens involved in this research.

5 Conclusions

As promoted by WHO, years lived in old age must not to be years of suffering and anguish, incapacities and dependencies, but rather years of meaning and quality of life, triggering processes that practice the available resources (individual and collective) in the redefinition of priorities, compensation of disabilities, adaptation to new situations, enabling the elderly to qualify, even when they have serious health problems enabling older people to remain a resource to their families, communities and economies [22].

We believe that the innovative work developed by Fonseca et al. [10, 16], and evaluated in this proposal on a different population sample, corroborate the recommendations in [10], i.e., the use of this type of Core Sets delineated the assess to the nursing care needs and/or to the outcomes of nursing interventions of citizens aged 65 years old or older, which will be also an ongoing process that will lead to the promotion of an Healthy Ageing and functional ability, which is stated as one of the frameworks promoted by WHO.

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Conflict of Interest. The authors declare that they have no conflicts of interest.

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